

The Modified Depression Scale (MDS): A Brief, No-Cost Assessment Tool to Estimate the Level of Depressive Symptoms in Students and Schools

Erin C. Dunn · Renee M. Johnson ·
Jennifer Greif Green

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Abstract Adolescent health researchers and practitioners are frequently interested in assessing depression as part of student screening and for school-wide prevention and intervention planning. However, this task is challenging given the lack of free, brief assessments of depressive symptoms in youth. This study evaluated the psychometric properties of an adapted version of the Modified Depression Scale (MDS). Data came from a school-based survey of 9th-12th graders in Boston ($N = 1,657$). We assessed internal consistency reliability and known-group validity, in addition to the feasibility of establishing a dichotomous cut-point to classify adolescents as having high versus low depressive symptoms. We also evaluated the validity of the adapted MDS as a school-wide measure. At the student level, the adapted MDS demonstrated acceptable internal consistency. Students engaging in risk behaviors (e.g., substance use) or who were victimized (e.g., bullied) had significantly higher depressive symptom scores. Students who endorsed four or five MDS symptoms often or always had a heightened risk of suicidal ideation, substance use, and failing grades when compared to students who

endorsed three or fewer symptoms often or always. At the school level, higher mean levels of depressive symptoms in a school were associated with higher mean levels of suicidal ideation and failing grades. Results of this study suggest that the adapted MDS is a promising measurement tool that could be useful to school-based professionals and researchers to evaluate depressive symptoms in adolescents and ascertain the prevalence of depressive symptoms in schools.

Keywords Depressive symptoms · Population health · Adolescents · Reliability · Validity

Introduction

Depression is one of the most common mental health problems among adolescents. As many as 30% of adolescents report experiencing symptoms of depression, such as feeling sad or hopeless, that interfere with their functioning (Eaton et al., 2008). Adolescents who experience depressive symptoms are more likely than their non-depressed peers to engage in a variety of risk behaviors, including unsafe sexual activity (Lehrer, Shrier, Gortmaker, & Buka, 2006), cigarette, alcohol, and drug use (Saban & Flisher, 2010), disordered eating (Fulkerson, Sherwood, Perry, Neumark-Sztainer, & Story, 2004), and suicidality (Hallfors et al., 2004). Students who experience depressive symptoms or engage in these risky behaviors are also more likely to have difficulties with learning and meeting the demands of the school environment (Eccles & Roeser, 2011; Hawkins, 1997). These findings suggest that preventing depression among students should be a priority for schools and local or state education authorities.

E. C. Dunn (✉)
Department of Society, Human Development, and Health,
Harvard School of Public Health, 677 Huntington Avenue,
Boston, MA 02115, USA
e-mail: erindunn@hsph.harvard.edu

R. M. Johnson
Department of Community Health Sciences, Boston University
School of Public Health, Boston, MA, USA
e-mail: rjohnson@bu.edu

J. G. Green
Special Education Program, Boston University
School of Education, Boston, MA, USA
e-mail: jgreen@bu.edu

Recognition of the link between depression, risk behaviors, and academic functioning has led to a growing interest in school-wide or population-based prevention and intervention activities focused on both educational and mental health outcomes among all students in a school (Jones, Brown, and Aber, 2011). These efforts are referred to as universal prevention in that they target the entire school population, rather than individual students, as in the case of a selected or indicated intervention (Mrazek & Haggerty, 1994). To track the progress of these school-wide prevention efforts, and identify the need for them in the first place, schools must monitor the level of depressive symptoms in the entire school (Dowdy, Ritchey, & Kamphaus, 2010). Tracking the prevalence of depressive symptoms among all students in a school is in keeping with a coordinated school health program, which emphasizes not only the provision of counseling, psychological, and social services, but also the assessment of mental, emotional, and social health problems in schools (McKenzie, Pinger, & Kotecki, 2008). Population-based assessments of whole school health are also important given that there is variation across schools in the prevalence of depression. For example, a study using data from the National Longitudinal Study of Adolescent Health found that approximately 5% of the variation in student's reports of depressive symptoms was due to the differences across schools (Dunn & Masyn, 2011). While this level of variation may seem small, when considered from a population-based perspective, it suggests that in addition to identifying students who experience depressive symptoms, districts will also benefit from identifying differential school need for interventions.

To engage in school-wide prevention and intervention activities, schools have traditionally used measures that focus on individual students. These include *diagnostic* assessments, which seek to determine the nature of individual student emotional problems, and *screening* assessments, which seek to identify individual students who are at-risk of emotional problems and who may therefore benefit from intervention (Glover & Albers, 2007). To complement these tools, schools also need population-based measures, which go beyond classification of individual students and allow schools to engage in: (a) *surveillance*, to monitor trends in levels of symptoms and associated risk factors among the school population (Dowdy et al., 2010); (b) *prediction*, to estimate schools at the greatest risk for poor outcomes; and (c) *evaluation*, to assess the efficacy of a given school-wide prevention or intervention program. For any measure to be practical for school-based researchers or school professionals, whether for use with individual students or entire schools, it must be inexpensive and easy to administer.

Currently, there are few measurement tools focusing on depressive symptoms that are ideally suited to researchers working with schools. As described in review articles (see for example Brooks & Kutcher, 2001; Myers & Winters, 2002), the most widely used self-report measures of child and adolescent depressive symptoms, such as the Beck Depression Inventory, include more than 20 items (and frequently more than 25), a length that discourages many school-based researchers and those interested in health behaviors from assessing depressive symptoms, altogether, particularly when resources are limited. Moreover, most instruments cannot be shortened or modified due to copyright restrictions and must be purchased from publishers, making them too costly for school-based researchers. At the same time, existing low-cost measures used for public health surveillance often fail to capture comprehensive information about depressive symptoms. For instance, the Youth Risk Behavior Surveillance System (YRBS) survey includes only one item to measure depressive symptoms (i.e., "During the past 12 months, did you ever feel so sad or hopeless almost every day for 2 weeks or more in a row that you stopped doing some usual activities?") (Centers for Disease Control and Prevention, 2005). Thus, there is a need for a brief and publicly available tool that: (a) could be easily administered to a large number of students; (b) provides information about the levels of depressive symptoms within a school; and (c) can be used for school-wide planning activities.

After reviewing the literature, we located the Modified Depression Scale (MDS)¹ in the Centers for Disease Control and Prevention's (CDC) compendium of assessment tools for youth violence research (Dahlberg, Toal, Swahn, & Behrens, 2005). The MDS is a self-administered six-item scale designed to assess the frequency of depressive symptoms (e.g., sadness, irritability, hopelessness, sleep disturbance, and concentration difficulties) among adolescents. It is based on the Roberts 26-item DSM Scale for Depression (Kelder et al., 2001; Roberts, Roberts, & Chen, 1997) and was adapted by Bosworth et al. (1996) and published in the compendium in its adapted form. The MDS is comparable in length to other measures, including the nine-item Patient Health Questionnaire (PHQ) that was also developed based on DSM criteria (Kroenke, Spitzer, & Williams, 2001). Although the PHQ assesses many of the same symptoms as the MDS, it has been primarily used with people older than age 18. Importantly, it does not assess irritability, which is one of the core features of depression in adolescents (American Psychiatric

¹ The CDC Compendium inaccurately attributed the MDS to Pamela Orpinas rather than Kris Bosworth (personal communication, Dr. Linda Dahlberg, Centers for Disease Control and Prevention, 9/21/2009).

Association, 2000). The MDS is neither a diagnostic nor a classification tool, but instead can be used to identify students at risk of depression.

Even though the MDS could be useful to school-based researchers and practitioners and has already been used in a number of studies, limited information exists about its psychometric properties. Studies using the MDS with adolescents in locations throughout the United States and Canada (Almeida et al., 2009; Bosworth et al., 1999; Edwards et al., 2006; Goldstein et al., 2007; Tandon & Solomon, 2009), including the World Health Organization-funded Health Behavior in School Age Children Study (Wang, Iannotti, Luk, & Nansel, 2010), have reported that the scale demonstrates good internal consistency reliability, with Cronbach's alpha coefficients of 0.74 or higher. However, detailed information about MDS reliability and validity is lacking. Specifically, little is known about the underlying factor structure of the MDS, whether MDS scores are associated with risk behaviors, and whether it is appropriate as a population-based tool.

This study was designed to examine the psychometric properties of an adapted version of the MDS in a school-based sample of adolescents for use as a tool to measure depressive symptoms at the level of students and schools. Our research questions were as follows: To what extent is the adapted MDS a reliable measure of depressive symptoms? To what degree are depressive symptoms scores, measured by the adapted MDS, associated with violence and educational outcomes that are known to be related to experiences of depression? Is it feasible to establish a dichotomous cut-point that could be used by schools to identify the percent of students with high levels of depressive symptoms? We explored this last question based on our interest in providing schools with a concise and easily understandable metric (i.e., a percentage) to describe the pervasiveness of depressive symptoms within a school. In evaluating the adapted MDS, we hypothesized that the scale would be a valid and reliable measure, capable of identifying both individual students and schools with high levels of depressive symptoms.

Methods

Subjects

Data came from the 2008 administration of the Boston Youth Survey (BYS), a biennial paper-and-pencil survey of high school students (9th–12th graders) in Boston Public Schools (BPS; Azrael et al., 2009). The BYS 2008 data collection instrument covered a range of topics (e.g., demographic characteristics, health behaviors, use of school and community resources, developmental assets, and risk factors) and had a particular emphasis on violence.

Thirty-two eligible public high schools within the BPS system were invited to participate. Additional schools that were considered ineligible for participation were those that served adults (i.e., “night schools”), students transitioning back to school after incarceration, students on suspension, and students who were primarily living outside of Boston (i.e., schools for children with special needs). Twenty-two eligible schools participated in the survey, for a school participation rate of 69%. Among schools considered eligible, we did not observe any statistically significant differences in key school indicators (e.g., drop-out rates, racial composition of students, and scores on standardized tests) between participating and non-participating schools.

To acquire a random sample of students within participating schools, we generated a list of unique humanities classrooms within each school. Classrooms were then stratified by grade and selected randomly for survey administration. Every student within the selected classrooms was invited to participate. Selection of classrooms continued until the total number of students to be surveyed ranged from 100 to 125 per school. In two schools with total enrollments close to 100, all classrooms in the school were invited to participate. An average of 85.4 students per school completed the survey (range: 40–132).

Instruments

Depressive Symptoms

We used an adapted version of the MDS to assess symptoms of depression. Students were asked to report the frequency of five symptoms in the past month: In the past month, how often... (a) “Were you very sad?”; (b) “Were you grouchy or irritable, or in a bad mood?”; (c) “Did you feel hopeless about the future?”; (d) “Did you sleep a lot more or less than usual?”; and (e) “Did you have difficulty concentrating on your school work?”. They used a five-point response scale: (1) never, (2) rarely, (3) sometimes, (4) often, and (5) always (The original option of “seldom” was replaced with “rarely” based on feedback from students who participated in pilot testing; they were unfamiliar with the word seldom.). The BYS research team omitted the sixth MDS item on appetite disturbance (“How often did you feel like not eating or eating more than usual?”) to conserve space, with the assumption that the item would be the least discriminating. We derived total scores by summing items among youth who had complete responses for all five items (range: 5–25). We also created a set of categorical variables denoting the percentages of youth who responded “often” or “always” to each survey item as a way of capturing more severe symptoms. In addition, we created school-level means on the adapted MDS to describe the average levels of depression within a school.

Related Behaviors and Indicators

In addition to the adapted MDS, we measured four indicators of youth well-being: (a) mental health and suicidality; (b) substance use; (c) victimization; and (d) school performance. To assess mental health and suicidality, students were asked whether they had, within the 12 months preceding survey administration, considered suicide, attempted suicide, cut or injured themselves, and visited a mental health professional. We assessed substance use by asking whether students used alcohol, tobacco, or marijuana in the past 30 days. We assessed victimization through a series of questions asking about being bullied, physical victimization by peers (e.g., punched and kicked by other kids), and family violence (e.g., punched, grabbed, by an adult in the home). School performance was assessed by asking respondents about their grades and truancy. Students who reported having “mostly D’s” or “mostly F’s” were classified as having failing grades and those who reported skipping school one or more times in the past month were classified as truant. Items related to suicidal ideation, suicide attempts, and substance use were adapted from the 2005 Youth Risk Behavioral Surveillance System questionnaire (Centers for Disease Control and Prevention, 2005). Items related to physical victimization from parents and peers were modified from the Conflict Tactics Scales (CTS Straus, 2007). The item on bullying was adapted from Rigby (1998). All other items were developed by the BYS research team. For school-level analyses, we created school-level means for each behavior or risk factor status by taking the average within each school.

Demographic Characteristics

The BYS also measured student’s grade in school, sex, Hispanic ethnicity, nativity (i.e., United States or foreign born), and race. We used a race/ethnicity variable with five levels: (a) Hispanic/Latino; (b) non-Hispanic, black/African American; (c) non-Hispanic, white; (d) non-Hispanic, Asian, and (e) Other, which includes those who were bi- or multiracial, Native Hawaiian or other Pacific Islander, American Indian, or Alaska Native. We also created school-level means for three of these variables, representing the proportion of students in the school who were female, non-Hispanic white, and foreign born.

Procedure

The paper-and-pencil survey was administered to students by trained staff between January and April of 2008. Prior to survey administration, we obtained passive consent from students’ parents. Survey administrators also read a statement on informed consent before each survey was

distributed. Students were given 50 min to complete the survey. Of the 2,725 students enrolled in the classrooms selected for participation, 1,878 completed a survey (response rate = 68.9%). Students who did not complete a survey either (a) chose not to participate ($n = 99$), (b) were not permitted by their parent to take the survey ($n = 24$), or (c) were absent from school on the day of survey administration ($n = 724$).

Data Analysis

Given that we were interested in evaluating the adapted MDS for the use as both an individual and school-level tool to evaluate levels of depressive symptoms, we conducted our analyses at both the individual and school levels. We began by conducting univariate analyses to describe the distribution of each item and the total MDS scale in this sample and among males and females. We calculated an intraclass correlation coefficient (ICC) to determine the proportion of variation in depressive symptoms that was due to difference across schools (rather than differences within schools, across students). Next, we conducted bivariate analyses to determine whether mean scores and proportions significantly varied across subgroups. To assess reliability, we conducted a categorical factor analysis, a type of analysis designed for ordinal data. In contrast to a continuous factor analysis, a categorical factor analysis does not require that data are continuous or that any distributional assumptions, such as normality, are met (Flora & Curran, 2004). We also used tetrachoric correlations, designed for categorical data, to examine the strength, direction, and magnitude of the associations between each MDS item. We then estimated the internal consistency reliability of the total MDS score by calculating Cronbach’s alpha coefficients for the total sample and separately for males and females. We also examined item-to-total score correlations, Cronbach’s alpha coefficients when specific items of the scale were deleted, and evaluated whether floor or ceiling effects were present (Terwee et al., 2007). We sought to provide information for known-group validity, a type of convergent construct validity (DeVellis, 2003), by comparing risk behavior and victimization status on the basis of MDS total scores, because youth who engage in risk behaviors or experience violence are more likely to have higher levels of depressive symptoms. We also examined, at the school level, the association between average MDS scores and the prevalence of engagement in risk behaviors. Based on our interest in determining whether the adapted MDS could be used as a binary indicator to differentiate youth with high levels of depressive symptoms from those with lower levels of depression or no symptoms, we created a set of binary variables indicating the number of youth who endorsed

“often” or “always” to each MDS item. We calculated odds ratios denoting the risk of experiencing each risk behavior or type of violence based on the number of items where youth reported high levels of depressive symptoms (e.g., endorsed often/always).

We conducted analyses with SAS version 9.2 ($\alpha < 0.05$ was the level of statistical significance). We stratified some analyses by sex based on well-documented sex differences in depression (Nolen-Hoeksema & Hilt, 2009) and restricted analyses to students who completed all five items on the adapted MDS scale. For individual-level descriptive statistics and tests of validity, in which standard errors were estimated, we accounted for the correlated nature of the observations or clustering of students within school by using generalized estimating equations (GEE) (Hanley, Negassa, de Edwardes, & Forrester, 2003). All individual-level tests of validity controlled for sex, age, and race/ethnicity. All school-level tests of validity were conducted using multiple linear regression models that were adjusted for the percentage of students in the school who were female and non-Hispanic white. We conducted analyses for the categorical factor analysis in Mplus version 6.1, which accounted for clustering of students in schools. Acceptable model fit was determined by a non-significant chi-square test, CFI and TLI values greater than 0.95, and RMSEA and SRMR values below 0.05 (Kline, 2010).

Results

Sample

To conduct psychometric analyses using a complete case analysis, we removed 221 youth (11.8%) missing data on any of the five MDS items, creating an analytic sample of 1,657. Missingness ranged from 10.3% on the sleep disturbance item to 8.8% on the sadness item. More than half of respondents in the analytic sample were female (54.0%). The sample was ethnically diverse, with 42.4% of students identifying as non-Hispanic black, 33.4% Hispanic/Latino, 9.3% non-Hispanic white, 8.1% Asian, and 6.7% Other or Multiracial. Respondents ranged in age from 13 to 19 years (8.0% were ≤ 14 years; 19.0% were 15; 27.7% were 16; 26.8% were 17; and 18.5% were ≥ 18 years) and were distributed almost equally across grades (24.4% from 9th grade; 27.2% from 10th grade; 26.8% from 11th grade; and 21.5% from 12th grade). Nearly one-third was born outside of the United States. Participants excluded from the analytic sample were not significantly different from those in the analytic sample with respect to race/ethnicity or age, though a greater percentage of excluded participants were male (58.4% vs. 46.1%, $z = 0.002$). With respect to the

school-level characteristics, the average school consisted of half males and females ($M = 52.61$; range: 22.0–69.0% female), a majority of students who were non-white ($M = 7.34$; range: 0–47.3% non-Hispanic white) and almost a third of students who were foreign born ($M = 31.08$; range: 10.3–61.8% foreign born).

Student-Level Descriptive Statistics

Table 1 describes the distribution of adapted MDS scores in the total sample and stratified by sex. Items corresponding to irritability and sleep problems had the highest mean values ($M = 2.9$ for each), and the item related to feeling hopeless had the lowest ($M = 2.1$). Females had significantly higher mean total scores ($z = 8.81$, $p < 0.0001$) as well as higher scores on each of the five items. Females more often reported experiencing MDS items often/always.

We examined the distribution of adapted MDS scores across groups expected to differ in their levels of depressive symptoms. Average MDS scores were significantly higher among older students (z -statistic = 2.58, $p = 0.01$) and differed by race/ethnicity (z -statistic = 0.06, $p = 0.02$). However, there were no statistically significant differences in MDS scores among those born in the United States compared with those born in another country (z -statistic = -1.11 , $p = 0.27$).

School-Level Descriptive Statistics

Figure 1 presents the distribution of MDS scores by schools. Mean depressive symptoms scores across schools ranged from 11.7 to 14.7, with the average depression score across schools being 13.2 (SE = 0.16, $p < 0.0001$). Although the variation within schools was greater than the variation across schools, there was significant between-school variation in depressive symptoms ($\sigma_{\text{between}}^2 = 0.30$, $p = 0.03$). Specifically, the ICC showed that 1.6% of the variability in depressive symptoms could be attributed to differences across schools.

Factor Structure and Correlations

The adapted MDS was unidimensional (*unrotated eigenvalues* = 2.96, 0.72); this one factor had good overall fit ($\chi^2 = 242.77$; CFI = 0.94; TLI = 0.88; RMSEA = 0.17; and SRMR = 0.05). The proportion of variance in each item (or observed variable) that was explained by the factors ranged from 40% for sleep problems to 57% for sadness. There were high and consistent factor loadings for all items (0.63–0.75). Factor loadings were highest for items measuring intrapsychic depressive symptoms, such

Table 1 Descriptive statistics for the modified depression scale (MDS) by each item and the total score

Survey item In the past 30 days, how often	Total sample (<i>n</i> = 1657)			Females (<i>n</i> = 891)			Males (<i>n</i> = 762)			z-statistic for means	z-statistic for proportions
	<i>M</i>	(SE)	% Reporting Often/ Always	<i>M</i>	(SE)	% Reporting Often/ Always	<i>M</i>	(SE)	% Reporting Often/ Always		
Were you very sad? (Sadness)	2.6	(0.04)	20.0	2.9	(0.03)	27.6	2.2	(0.05)	11.3	12.91***	9.10***
Were you grouchy, irritable, or in a bad mood? (Irritability)	2.9	(0.05)	26.7	3.1	(0.05)	33.5	2.7	(0.06)	18.6	7.91***	8.60***
Did you feel hopeless about the future? (Hopelessness)	2.1	(0.04)	15.6	2.3	(0.05)	18.7	2.0	(0.06)	11.9	3.94***	3.24**
Did you sleep a lot more or a lot less than usual? (Sleep problems)	2.9	(0.04)	31.6	3.0	(0.06)	36.0	2.7	(0.04)	26.4	4.65***	3.77***
Did you have difficulty concentrating on your school work? (Concentration difficulties)	2.8	(0.04)	28.5	2.9	(0.06)	30.9	2.7	(0.05)	25.6	3.39***	3.53***
Total score	13.3	(0.17)	55.9	14.2	(0.17)	62.6	12.2	(0.19)	47.8	8.81***	6.89***

Cell entries are estimates of the mean, adjusted standard errors (adjusted for clustering), the percentage of respondents who reported often/always for each item and on every item (total score), as well as *t* statistics and χ^2 statistics for analyses comparing females and males. Results derived from generalized estimating equations (GEE)

~ *p* < 0.10; * *p* < 0.05; ** *p* < 0.001; *** *p* < 0.0001

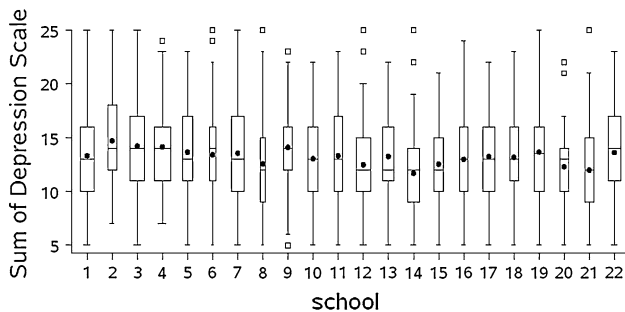


Fig. 1 Distribution of MDS scores within and between schools (*n* = 22). The figure presents *box plots* depicting the distribution of MDS scores within each school and across all of the schools in the sample. *Dots* represent the mean adapted MDS scores within the school. *Horizontal lines* in the *boxes*, for each school, represent the first quartile, the median, and the third quartile of the distribution of adapted MDS total scores within each school. *Vertical lines* depict the range of total scores (minimum to maximum values) in each school. *Squares* depict outlying values

as sadness and irritability (0.75 and 0.73, respectively), and lowest for items assessing somatic symptoms, such as sleep disturbance and difficulty with concentration (0.63 and 0.66, respectively). Tetrachoric correlation coefficients indicated that MDS items were moderately and significantly correlated (*r* = 0.40–0.62). The intrapsychic symptoms of sadness, irritability, and hopelessness were most strongly correlated with one another. Sleep disturbance had the weakest correlation with other items (*r* = 0.44–0.40).

Floor or Ceiling Effects

The adapted MDS did not have floor or ceiling effects; that is, no more than 15% of the sample had the lowest or highest possible score. Total adapted MDS scores were normally distributed, with 4.5% of students having the lowest score (5) and 0.9% having the highest (25).

Internal Consistency Reliability

The adapted MDS displayed acceptable internal consistency reliability in the total sample (Cronbach's α = 0.79; Table 2). The Cronbach's alpha coefficient did not increase with the deletion of any individual items; it decreased the least with the deletion of sleep problems. Findings from sex-stratified analyses were similar to the total sample. However, for females, the items corresponding to sadness and hopelessness correlated most strongly with the total score, whereas hopelessness and irritability were most strongly correlated with the total score for males.

Known-Group Validity

Table 3 presents the result of analyses examining the association between victimization and risk behaviors with MDS scores. Odd ratios (OR) refer to the odds of each risk status for each one unit difference in the MDS symptom total score. The adapted MDS significantly differentiated

Table 2 Internal consistency reliability values of the modified depression scale (MDS) in the BYS sample

Survey item	Total sample (<i>n</i> = 1,657)			Females (<i>n</i> = 891)			Males (<i>n</i> = 762)		
	<i>r</i> to total score	α	α if item deleted	<i>r</i> to total score	α	α if item deleted	<i>r</i> to total score	α	α if item deleted
Sadness	0.60		0.74	0.60		0.72	0.57		0.74
Irritability	0.59		0.74	0.57		0.73	0.58		0.74
Hopelessness	0.59		0.74	0.59		0.72	0.59		0.73
Sleep problems	0.51		0.77	0.47		0.76	0.52		0.76
Concentration difficulties	0.55		0.76	0.55		0.74	0.55		0.75
Total score		0.79			0.78			0.78	

Cell entries are raw Pearson's *r* correlation coefficients between individual items and the total MDS score, Cronbach's alpha coefficient, and Cronbach's alpha coefficient if the individual items were deleted

youth who had engaged in risk behavior or who had been victimized from those who had not (OR ranged from 1.07 to 1.31, all *p* < 0.001). For example, the odds of considering suicide were 1.31 times higher for each one unit difference in the adapted MDS score. The strongest association was between the MDS and reports of suicidal ideation.

Table 4 presents the results of the school-level analyses examining the association between school-level depressive symptoms and school-level risk behavior problems. The adapted MDS significantly differentiated schools with a high prevalence of students who considered suicide and received failing grades. Specifically, we found that each

one unit difference in the average MDS score within a school was associated, on average, with a 5.3% difference in the proportion of students who had received failing grades and an 8.2% difference in the proportion of students who had considered suicide. However, MDS scores were not significantly associated with any other school-level risk factors.

In examining frequent depressive symptoms, we found that 55.9% reported experiencing at least one MDS symptom often or always (last row of Table 1). Moreover, 34.1% of the total sample reported experiencing two or more symptoms at this frequency, 18.3% reported three or

Table 3 Results of student-level analyses estimating the log odds of behavior/victimization status based on MDS total score

	OR	95% CI
Mental health and suicidality		
Considered suicide (<i>n</i> = 197)	1.31	1.26–1.37
Attempted suicide (<i>n</i> = 66)	1.27	1.20–1.34
Cut/injured self (<i>n</i> = 141)	1.21	1.16–1.25
Visited professional (<i>n</i> = 382)	1.14	1.11–1.17
Substance use		
Drank alcohol (<i>n</i> = 621)	1.10	1.07–1.13
Used tobacco (<i>n</i> = 205)	1.16	1.10–1.21
Used marijuana (<i>n</i> = 320)	1.07	1.03–1.11
Victimization		
Bullied by peers (<i>n</i> = 677)	1.12	1.09–1.15
Victimized by peers (<i>n</i> = 370)	1.10	1.07–1.12
Experienced family violence (<i>n</i> = 263)	1.14	1.09–1.20
School performance		
Failing grades (<i>n</i> = 243)	1.11	1.08–1.15
Truant (<i>n</i> = 588)	1.08	1.05–1.11

Cell entries are adjusted odds ratios (adjusted for clustering) and 95% CI (confidence interval). All OR were statistically significant at *p* < 0.0001. Cell sizes across analyses may not sum to total *N* due to missing data. All analyses controlled for sex, age, and race/ethnicity

Table 4 Results of school-level analyses estimating the prevalence of behavior/victimization status based on MDS total score

	Beta (SE)	<i>p</i> value
Mental health and suicidality		
Considered suicide (12.2)	8.21 (2.75)	0.01
Attempted suicide (4.6)	5.50 (3.63)	0.15
Cut/injured self (8.7)	6.45 (3.95)	0.12
Visited professional (24.3)	1.11 (1.73)	0.53
Substance use		
Drank alcohol (38.9)	−0.16 (1.45)	0.91
Used tobacco (13.0)	0.35 (2.06)	0.87
Used marijuana (20.4)	−0.73 (1.85)	0.70
Victimization		
Bullied by peers (41.5)	1.59 (1.39)	0.27
Victimized by peers (23.8)	0.72 (1.85)	0.70
Experienced family violence (16.5)	2.17 (2.50)	0.40
School performance		
Failing grades (16.3)	5.30 (1.56)	0.01
Truant (38.0)	0.76 (1.11)	0.50

Cell entries are beta-coefficients and standard errors (SE). Parentheses next to each risk factor indicator refer to the average level of each indicator across all schools (i.e., the average school-wide percent). All analyses controlled for percentage of students in the school who were female and white

more symptoms, 9.6% reported four or more symptoms, and 4.4% endorsed all five MDS items often/always. Figure 2 presents the results of analyses investigating whether there was an appropriate cut-point to differentiate youth with high depressive symptoms from all others. The figure is organized according to the number of items youth endorsed as often/always, with adjusted odds ratios that compare the log odds of each risk behavior among those with a score of one or more, two or more, three or more, four or more, or five on the MDS (versus youth with fewer symptoms). Results of these analyses suggest that several outcomes increased considerably among youth who endorsed four or five (vs. three or fewer) MDS symptoms: considered suicide, attempted suicide, cut/injured self, visited a mental health professional, drank alcohol, used tobacco, and reported failing grades. However, this pattern was not consistent across other outcomes, which either increased linearly and did not have a clear cut-point (e.g., family violence) or were unrelated to MDS scores (e.g., used marijuana, bullied, and truant). The odds of experiencing each outcome increased from 1.90-fold to 5.53-fold among those reporting one or more symptoms of depression (compared to those reporting no symptoms) to between 2.12-fold and 9.54-fold among those reporting all five symptoms (compared with those reporting fewer).

Discussion

In response to the need for a brief, reliable, and inexpensive measure of depressive symptoms for researchers in schools, we set out to assess the psychometric properties of an adapted Modified Depression Scale (MDS) for use with adolescents. Results demonstrate that the adapted MDS has good reliability and validity. Consistent with other studies of the MDS (Bosworth et al., 1999; Edwards et al., 2006; Goldstein et al., 2007), we found that—in a school-based sample of Boston youth—the adapted MDS had acceptable internal consistency reliability ($\alpha = 0.79$). It also did not have floor or ceiling effects. We extend previous work on the MDS by showing that the adapted version had a clear unidimensional structure with highest factor loadings for items that characterize intrapsychic symptoms of depression (sadness, irritability, and hopelessness). This is in keeping with current thinking regarding the nosology of depression in youth, as the DSM-IV identifies sadness and irritability (in children and adolescents) as core symptoms of depression (American Psychiatric Association, 2000). By comparison, the somatic symptoms measured by the MDS (sleep problems and concentration difficulties) had lower factor loadings and reduced correlations with other items. Although somatic symptoms are listed as indicators of depression in the DSM-IV, they can also be markers of a

number of other psychological problems (e.g., anxiety, attention-deficit/hyperactivity disorder, and substance use); it is therefore not surprising that they were less strongly associated with the total score.

This study also shows that the adapted MDS demonstrated strong known-group validity, as it differentiated both youth and schools on the basis of several risk factors. Consistent with current literature on risk factors associated with depression, those with higher MDS total scores were more likely to be female (Nolen-Hoeksema & Hilt, 2009), have experienced peer victimization and bullying (Klomek, Marrocco, Kleinman, Schonfeld, & Gould, 2007), have recently used alcohol, tobacco, and marijuana (Armstrong & Costello, 2002; Kandel et al., 1997), and display other indicators of psychological distress, including suicidal ideation, self-injury, and receipt of psychological services (Alexandre, Dowling, Stephens, Laris, & Rely, 2008; Fergusson, Beautrais, & Horwood, 2003; Hallfors et al., 2004; Osman, Kopper, Barrios, Gutierrez, & Bagge, 2004). Odds ratios for these effects were high (range: 1.07–1.31), especially as they represented the odds of each risk status for just a *one unit difference* in the MDS score. These findings suggest that the adapted MDS is useful for predictive purposes in that it can effectively serve as an indicator of depressive symptoms that increase adolescents' risk for a number of poor psychosocial outcomes. Based on these results, school researchers should coadminister the MDS with other brief measures of psychosocial functioning in an effort to understand predictors or correlates of depressive symptoms at both the level of students and schools (Dowdy et al., 2010). These types of comprehensive measurement tools have been used with other mental health outcomes in youth, including eating disorders (Shisslak et al., 1999), and would enable schools to understand school-wide levels of these co-occurring experiences.

We also found that schools with high average MDS scores also have, on average, higher levels of students who considered suicide and received failing grades. These findings suggest that the MDS may be able to detect not just individual students in need, but also schools in need of global prevention and intervention services. The finding that the adapted MDS was more strongly associated with psychosocial risk factors at the student level, compared to risk factors at the school level, suggests that the MDS may be a better predictor of student need rather than school need. However, our study may not have been powered to the same degree at both levels of analysis (i.e., at the student level and school level). That is, although we included a large number of students in this sample, we included a relatively small number of schools and only had modest levels of between-school variation in depressive symptoms (even though this level was statistically significant). Thus, the finding that school levels of considering suicide and

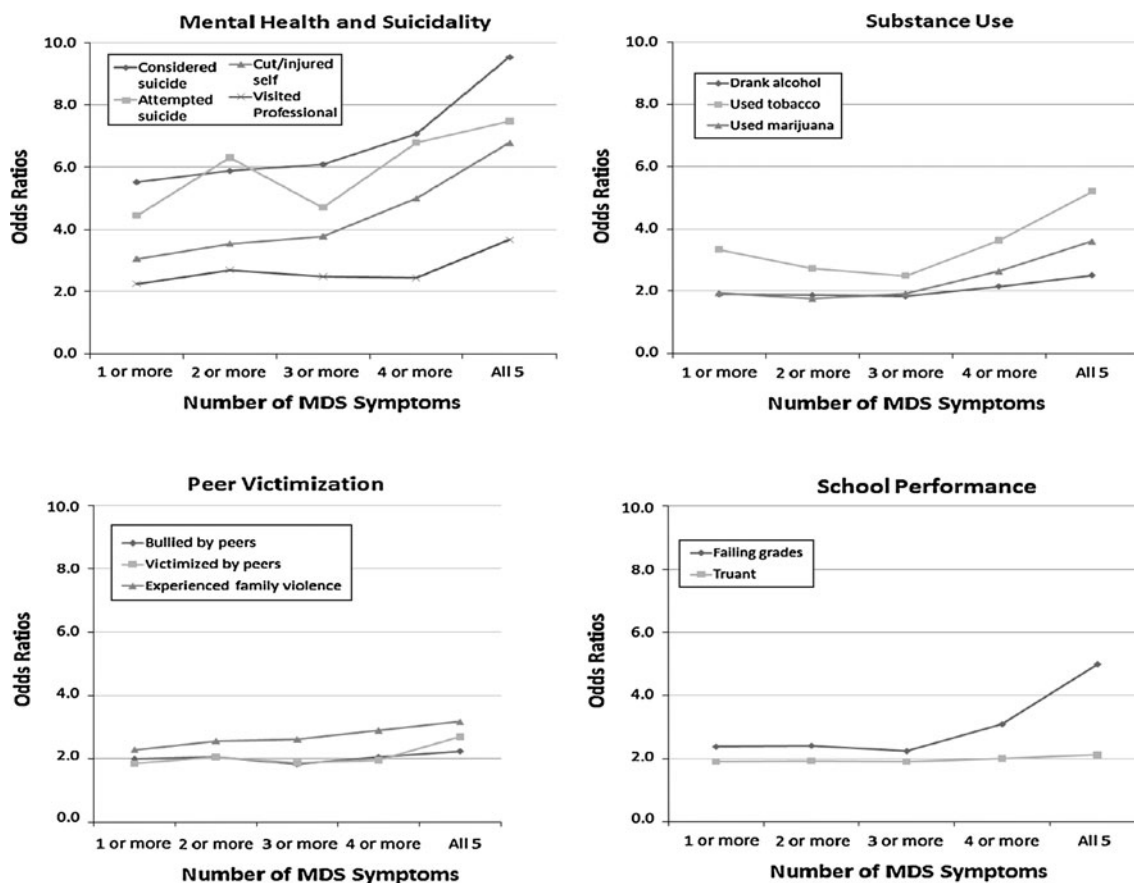


Fig. 2 Results of generalized estimating equation (GEE) analyses estimating the log odds of each behavior/victimization status by MDS category score. Cell entries are adjusted odds ratios (adjusted for clustering) and 95% CI. All OR were statistically significant at $P < 0.0001$. The referent group for each comparison (i.e. OR) is youth with a fewer number of symptoms (e.g., 1 or more symptoms

compared to zero; 2 or more symptoms compared to zero or 1 symptoms). Analyses controlled for sex, age, and race/ethnicity. Sample sizes for each MDS score were: $n = 926$ endorsed 1 or more MDS symptoms, $n = 565$ endorsed 2 or more MDS symptoms, $n = 304$ endorsed 3 or more MDS symptoms, $n = 159$ endorsed 4 or more MDS symptoms, $n = 73$ endorsed all 5 symptoms often/always

receiving failing grades were associated with school levels of depressive symptoms suggests that these associations are quite robust. Indeed, just a one-unit difference in the average MDS score within a school was associated, on average, with a 5.3% difference in the proportion of students who had received failing grades and an 8.2% difference in the proportion of students who had considered suicide. Future studies that include a larger number of schools will be better able to detect more moderate associations.

Further research is also needed on the most effective application of the MDS to estimating school-wide risk of depression and informing prevention and intervention planning. More detailed school-level analyses of the association between the related risk factors and MDS scores, as well as school-level demographic features (e.g., school demographic composition; teacher training and tenure) could identify important between-school differences that have relevance to program implementation. These types of

analyses will be important for determining the applicability of the MDS for mental health services planning and determining the schools most in need of prevention and intervention efforts.

Analyses to determine whether the adapted MDS could be used with a dichotomous cut-point suggest that adolescents who experienced any of the MDS symptoms often/always had significantly increased risk of most of the psychosocial outcomes we considered. Endorsing four or five MDS symptoms heightened the risk of suicidality, substance use, and failing grades. These findings were not surprising, as the association between these risk factors and depression is well established. In contrast, the cut-point did not distinguish between adolescents with high and low levels of depressive symptoms on more distal outcomes (e.g., peer victimization, truancy). Rather than being linked directly, these indicators may represent mediating pathways (e.g., depressive symptoms cause weight gain, which in turn leads to being bullied by peers). We also found that

the proportion of students endorsing four or more depressive symptoms often/always (9.6%) was consistent with existing epidemiological research on the prevalence of major depression in adolescents (Green, McGinnity, Meltzer, Ford, & Goodman, 2005; Merikangas et al., 2010). Based on these findings, we think researchers interested in the MDS as a binary indicator (depressed/not depressed) may find the ≥ 4 cut-point useful.

Limitations and Future Directions

This study had several limitations. First, the BYS research team omitted the item measuring appetite disturbance from the MDS. This was done to shorten the data collection instrument so as to reduce respondent burden and keep the instrument short enough to be completed within a class period. This decision was made on the basis of familiarity with the instrument, rather than empirical and psychometric data. Given our finding that the MDS somatic items (i.e., sleep problems and concentration difficulties) had lower factor loadings than the intrapsychic items, and consistent with other studies (Edwards et al., 2006), we suspect that appetite disturbance would have also had a reduced association with other items and thus provided less information about depressive symptoms. Future research should examine this. Second, all measures included in this study are based on self-report of adolescents, and it is possible that the associations between the MDS and other risk factors were artificially inflated by common method variance. Although it is often difficult for schools to collect data from multiple informants (e.g., teachers, parents), conducting a multi-informant analysis could provide additional information about the validity of the MDS. Third, although our results indicate that the adapted MDS is strongly associated with risk factors related to depression, without a coadministered “gold standard” assessment of depressive symptoms, it is difficult to fully ascertain the validity of the MDS. Future research on the MDS should include a coadministered depressive symptoms measure.

Conclusions

Results of this study provide good preliminary evidence to document the reliability and validity of the MDS, which is increasingly being used in large school-based studies to measure depressive symptoms among youth (e.g., Wang et al., 2010). The MDS fills an important need for school researchers as a brief, inexpensive, and publicly available tool for evaluating depressive symptoms in adolescents and ascertaining the prevalence of depressive symptoms in schools. By surveying student depressive symptoms, schools can determine the level of need of their students,

detect trends over time, and select appropriate prevention and intervention programs. This type of surveillance is critical for reducing depressive symptoms, monitoring associated risk behaviors, and ultimately improving the health and well-being of school children. The brevity of the MDS, combined with its psychometric properties, makes it well suited for survey activities that are limited in their resources or space. The MDS is therefore a very good option for both schools who want to conduct epidemiological activities and school-based researchers interested in incorporating a measure of depression into their study.

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